Some details on two-phase variances

Thomas Lumley

May 15, 2006

This document explains the computation of variances for totals in two-phase designs. Variances for other statistics are computed by the delta-method from the variance of the total of the estimating functions.

The variance formulas come from conditioning on the sample selected in the first phase

$$\mathrm{var}[\hat{T}] = E\left[\mathrm{var}\left[\hat{T}|\text{phase 1}\right]\right] + \mathrm{var}\left[E\left[\hat{T}|\text{phase 1}\right]\right]$$

The first term is estimated by the variance of $\hat{T}$ considering the phase one sample as the fixed population, and so uses the same computations as any single-phase design. The second term is the variance of $\hat{T}$ if complete data were available for the phase-one sample. This takes a little more work.

The variance computations for a stratified, clustered, multistage design involve recursively computing a within-stratum variance for the total over sampling units at the next stage. That is, we want to compute

$$s^2 = \frac{1}{n-1}\sum_{i=1}^{n}(X_i - \bar{X})$$

where X_i are π-expanded observations, perhaps summed over sampling units. A natural estimator of s^2 when only some observations are present in the phase-two sample is

$$\hat{s}^2 = \frac{1}{n-1}\sum_{i=1}^{n}\frac{R_i}{\pi_i}(X_i - \hat{\bar{X}})$$

where π_i is the probability that X_i is available and R_i is the indicator that X_i is available. We also need an estimator for $\bar{X}$, and a natural one is

$$\hat{\bar{X}} = \frac{1}{n}\sum_{i=1}^{n}\frac{R_i}{\pi_i}X_i$$

This is not an unbiased estimator of s^2 unless $\hat{\bar{X}} = \bar{X}$, but the bias is of order $O(n_2^{-1})$ where $n_2 = \sum_i R_i$ is the number of phase-two observations.

If the phase-one design involves only a single stage of sampling then X_i is Y_i/p_i, where Y_i is the observed value and p_i is the phase-one sampling probability. For multistage phase-one designs (not yet implemented) X_i will be more complicated, but still feasible to automate.

This example shows the unbiased phase-one estimate (from Takahiro Tsuchiya) and the estimate I use, in a situation where the phase two sample is quite small.

First we read the data

```
rei<-read.table(textConnection(
"  id   N n.a h n.ah n.h   sub  y
1   1 300  20 1   12   5  TRUE  1
2   2 300  20 1   12   5  TRUE  2
3   3 300  20 1   12   5  TRUE  3
4   4 300  20 1   12   5  TRUE  4
5   5 300  20 1   12   5  TRUE  5
6   6 300  20 1   12   5 FALSE NA
7   7 300  20 1   12   5 FALSE NA
8   8 300  20 1   12   5 FALSE NA
9   9 300  20 1   12   5 FALSE NA
10 10 300  20 1   12   5 FALSE NA
11 11 300  20 1   12   5 FALSE NA
12 12 300  20 1   12   5 FALSE NA
13 13 300  20 2    8   3  TRUE  6
14 14 300  20 2    8   3  TRUE  7
15 15 300  20 2    8   3  TRUE  8
16 16 300  20 2    8   3 FALSE NA
17 17 300  20 2    8   3 FALSE NA
18 18 300  20 2    8   3 FALSE NA
19 19 300  20 2    8   3 FALSE NA
20 20 300  20 2    8   3 FALSE NA
"), header=TRUE)
```

Now, construct a two-phase design object and compute the total of y

```
> des.rei <- twophase(id = list(~id, ~id), strata = list(NULL,
+     ~h), fpc = list(~N, NULL), subset = ~sub, data = rei)
> tot <- svytotal(~y, des.rei)
```

The unbiased estimator is given by equation 9.4.14 of Särndal, Swensson, & Wretman.

```
> rei$w.ah <- rei$n.ah/rei$n.a
> a.rei <- aggregate(rei, by = list(rei$h), mean, na.rm = TRUE)
> a.rei$S.ysh <- tapply(rei$y, rei$h, var, na.rm = TRUE)
> a.rei$y.u <- sum(a.rei$w.ah * a.rei$y)
> a.rei$f <- with(a.rei, n.a/N)
> a.rei$delta.h <- with(a.rei, (1/n.h) * (n.a - n.ah)/(n.a - 1))
> Vphase1 <- with(a.rei, sum(N * N * ((1 - f)/n.a) * (w.ah * (1 -
+     delta.h) * S.ysh + ((n.a)/(n.a - 1)) * w.ah * (y - y.u)^2)))
```

The phase-two contributions (not shown) are identical. The phase-one contributions are quite close

```
> Vphase1
```

```
[1] 24072.63
```

```
> attr(vcov(tot), "phases")$phase1
```

```
         [,1]
[1,] 23461.05
```